

Addiction History Associates with the Propensity to Form Habits

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Abstract

■ Learned habitual responses to environmental stimuli allow efficient interaction with the environment, freeing cognitive resources for more demanding tasks. However, when the outcome of such actions is no longer a desired goal, established stimulus–response (S-R) associations or habits must be overcome. Among people with substance use disorders (SUDs), difficulty in overcoming habitual responses to stimuli associated with their addiction in favor of new, goal-directed behaviors contributes to relapse. Animal models of habit learning demonstrate that chronic self-administration of drugs of abuse promotes habitual responding beyond the domain of compulsive drug seeking. However, whether a similar propensity toward domain-general habitual responding occurs in humans with SUDs has remained unclear. To address this question, we used a visuomotor S-R learning and relearning task, the Hidden

Association between Images Task, which employs abstract visual stimuli and manual responses. This task allows us to measure new S-R association learning and well-learned S-R association execution and includes a response contingency change manipulation to quantify the degree to which responding is habit-based, rather than goal-directed. We find that people with SUDs learn new S-R associations as well as healthy control participants do. Moreover, people with an SUD history slightly outperform controls in S-R execution. In contrast, people with SUDs are specifically impaired in overcoming well-learned S-R associations; those with SUDs make a significantly greater proportion of perseverative errors during well-learned S-R replacement, indicating the more habitual nature of their responses. Thus, with equivalent training and practice, people with SUDs appear to show enhanced domain-general habit formation. ■

INTRODUCTION

Learned habitual responses to stimuli allow efficient navigation of daily life by allocating cognitive resources toward processes such as cognitive control, which enables flexible behavioral. However, when the outcome of such habitual actions is no longer a desirable goal, established stimulus–response (S-R) associations must be overcome. A definitive behavior of addiction is continued drug use despite serious negative consequences of such use. In essence, although the outcome of drug seeking and/or consumption is reduced in value from mostly positive to mixed or largely negative, these actions persist and can be potently triggered by drug-associated cues. As such, addiction may be partially described as an initially goal-directed behavior that becomes a habit-based process as a consequence of reinforcement learning during repeated drug use (Belin, Belin-Rauscent, Murray, & Everitt, 2013; Everitt & Robbins, 2005, 2013; Balleine & O’Doherty, 2010; Belin, Jonkman, Dickinson, Robbins, & Everitt, 2009). Despite the clinical importance of understanding the maladaptively rigid behaviors that characterize substance use disorders (SUDs), investigation of behavioral rigidity in SUDs has been limited to date.

Data from animal models show that extended cocaine (Zapata, Minney, & Shippenberg, 2010; Belin & Everitt,

2008) or alcohol use (Corbit, Nie, & Janak, 2012; Dickinson, Wood, & Smith, 2002) promotes habitual behavior, suggesting that chronic exposure to drugs of abuse potentiates habitual responding more generally. In contrast to what is known in animals, relatively little is known about habit learning in humans, or whether addiction is associated with altered general capacity to learn or replace S-R associations, that is, to form or break habits. Either enhanced habit formation or impaired ability to overcome habits could theoretically contribute to addiction.

Animal studies of S-R learning are typically limited to simple one-to-one mapping of stimuli onto response options, and although such designs have been used with humans (Toni, Ramnani, Josephs, Ashburner, & Passingham, 2001; Deiber et al., 1997), people learn such associations rapidly, limiting their use for examining learning over time and measuring transitions between goal-directed and habitual response selection. Moreover, habitual responding in animals is typically tested via outcome devaluation, with continued responding for a devalued outcome taken to indicate habitual responding (Dickinson, 1985). Outcome devaluation studies in humans replicate animal studies of simple S-R learning tasks (de Wit, Corlett, Aitken, Dickinson, & Fletcher, 2009; Tricomi, Balleine, & O’Doherty, 2009; Valentin, Dickinson, & O’Doherty, 2007), but such designs have substantial methodological limitations in humans. In particular, it is very difficult to identify multiple primary

reinforcers equated for value across individuals that may then be devalued according to traditional animal paradigms. This difficulty precludes their use with special populations, which pose other recruitment challenges, and renders them ill-suited to multisession tests of interventions to reduce habitual responding. Furthermore, these paradigms lack ecological validity for modeling stimulus-driven (i.e., habit-based) actions in the “real world.” During daily life, it is less often the case that the outcome of a no-longer adaptive action loses value; rather, it is that the outcome itself changes to one that is less (or un-) desirable. For example, the cue of walking into a darkened room will often trigger the automatic action of flipping the light switch. During a power outage, this action will yield no positive outcome, and yet this automatic action will persist despite its known lack of utility. In this case, the former outcome of this action (illumination of the darkened room) retains its value, but that is simply no longer the outcome of flipping the light switch. Likewise, in the case of maladaptive habit-based actions, such as compulsive drug use during a binge, the same action (e.g., lighting and smoking from a crack pipe) will no longer yield the initial outcome, a euphoric “high,” instead producing agitation and paranoid delusions. Again, the former outcome of this action (a euphoric high state) retains its value, but it is no longer the outcome of the action. Most work in humans to date has overcome these obstacles by instead employing probabilistic learning tasks (Dolan & Dayan, 2013). However, although such paradigms are useful in investigating the ability to flexibly adapt to dynamic response contingencies, these paradigms cannot promote enduring habitual responses to stimuli. Our task, although simplistic, provides a useful laboratory-based model of these sort of S-R outcome contingency changes that are a natural part of human life.

Few studies to date have investigated the relationship between habitual behavior and drugs of abuse in humans. First, adult smokers demonstrate both goal-directed and habitual responding for natural rewards and cigarettes, dependent on age, smoking habit severity, and impulsiveness (Hogarth, Chase, & Baess, 2012; Hogarth & Chase, 2011). Second, a recent neuroimaging study in alcohol-dependent patients, including those concurrently using psychoactive medications for depression and anxiety disorders, found preferential habit-based responding during task performance at the expense of goal-directed behavior (Sjoerds et al., 2013); however, confounding factors preclude strong linkage between habitual responding and alcohol use disorders.

Finally, no published work to date has investigated the transition between goal-directed and habitual response selection during S-R formation or the replacement of S-R associations in people with SUDs. To address this knowledge gap, we compared S-R association learning and replacement between healthy adults with no SUD history and currently abstinent people with a lifetime SUD diagnosis (Table 1). We predicted that an SUD history would associate with enhanced capacity to acquire new S-R asso-

ciations and impaired ability to replace established responses, with a specific increase in perseverative responding when attempting to change established S-R associations. To test these ideas, we employed the Hidden Association between Images Task (HABIT; Figure 1), a visuomotor S-R learning and relearning task with abstract visual stimuli and manual responses. As the behavioral data were not normally distributed, we applied generalized linear mixed-effects models (GLMMs) to characterize the change in behavioral performance over time, evaluating whether SUD status uniquely accounts for significant variability in learning and/or relearning trajectories across individuals.

METHODS

Participants

A total of 62 participants were recruited from the University of North Carolina at Chapel Hill (UNC) and the surrounding community via advertisements. Participants were recruited into two groups, based on whether they did ($n = 22$ SUD) or did not ($n = 40$ control; Ctrl) meet DSM-IV criteria for past drug or alcohol dependence in a structured clinical interview ($n = 7$ alcohol, $n = 4$ opiates, $n = 11$ stimulants, of which $n = 13$ were polysubstance abusers; Sheehan et al., 1998). SUD participants self-reported a minimum of 2 weeks of abstinence at the time of recruitment ($M = 2 \pm 2.5$ years). All participants were healthy individuals 18–40 years old with no known history of neurological disorders, no current psychiatric diagnoses ($n = 5$ SUDs met criteria for past depression) or psychoactive drug or medication use (excluding nicotine and caffeine), and reported an IQ within the normal range (≥ 80). Participants were screened for psychoactive drug use (Biotechnostix, Inc., Markham, Ontario), including alcohol (FC-10, Lifeloc, Inc., Wheat Ridge, CO) in each session. Thirteen additional participants were recruited but failed to complete the training session (see “Behavioral Task”), and another 12 participants failed to return for or complete the testing session. As expected, the SUD and control groups differed significantly in terms of substance use, with higher scores on all measures, including family history of alcohol abuse in the SUD group (Table 1). The SUD and control groups did not differ significantly in terms of education, socioeconomic status, gender, or ethnicity but did differ in terms of age and estimated IQ, with significantly lower average IQ and higher average age for the SUD group relative to the Ctrl group (Table 1); to control for these differences, age and IQ were included as covariates in all analyses. Each participant provided written informed consent as approved by the UNC Office of Human Research Ethics.

General Procedure

Individuals participated in two sessions, with at least one night’s sleep between the first and second sessions. Participants were paid for their involvement, including

Table 1. Sample Demographics and Psychometric Data

	<i>Ctrl (n = 40)</i>	<i>SUD (n = 22)</i>	<i>t(60)</i>	<i>p</i>
<i>Demographics</i>				
Age (years)	24 ± 6	29 ± 6	-2.65	.01
SILS (calculated) IQ	105 ± 6	99 ± 6	4.32	<.001
Education (years)	15 ± 2	15 ± 2	0.40	.69
SES	46 ± 9	41 ± 12	1.51	.14
Gender (% female)	50	16		<i>ns</i> ^a
Ethnicity (% non-White)	27	43		.28 ^b
<i>Substance Use Related</i>				
AUDIT Total	4 ± 3	23 ± 10	-8.60	<.001 ^c
AUDIT Consumption	3 ± 2	8 ± 3	-7.28	<.001
AUDIT Dependence	0.08 ± 0.35	6 ± 5	-8.31	<.001 ^c
AUDIT Harm	0.78 ± 1.33	8 ± 6	-7.19	<.001 ^c
DAST	1 ± 1	17 ± 7	-10.79	<.001 ^c
DUSI-I (%)	0.10 ± 0.12	0.80 ± 0.18	-16.25	<.001 ^c
FTQ density (%)	0.16 ± 0.22	0.41 ± 0.23	-4.28	<.001
<i>Psychometric</i>				
BDI	3 ± 4	6 ± 6	-1.69	.10 ^c
BIS Total	55 ± 8	68 ± 15	-3.61	.001 ^c
BIS Attention	14 ± 3	18 ± 5	-2.55	.01
BIS Motor	21 ± 3	25 ± 6	-2.73	.01 ^c
BIS Nonplanning	20 ± 4	25 ± 6	-4.31	<.001 ^c
LOC	10 ± 3	8 ± 4	1.99	.051
MMPI-Antisocial Practices Scale	6 ± 3	10 ± 5	-3.57	.001 ^c
STAI Total	60 ± 15	67 ± 15	-1.63	.11
STAI-State Anxiety	27 ± 7	29 ± 7	-0.96	.34
STAI-Trait Anxiety	33 ± 8	37 ± 9	-2.02	.048
TAF Total	17 ± 13	19 ± 14	-0.43	.67
TAF Moral	16 ± 11	15 ± 10	0.40	.69
TAF Self	1.2 ± 2.1	3.0 ± 3.6	-2.15	.03 ^c
TAF Others	0.6 ± 1.6	1.4 ± 3.1	-1.14	.26 ^c

Values are reported as mean ± standard deviation. Reported *p* values reflect the results of unpaired two-tailed comparison between groups. SUD = history of SUD participant; Ctrl = control subject; IQ = intelligence quotient; SES = socioeconomic status; AUDIT = Alcohol Use Disorders Identification Test; DAST = Drug Abuse Screening Test; DUSI-I = Drug Use Screening Inventory, Domain I; FTQ = Family Tree Questionnaire; BDI = Beck Depression Index; BIS = Barratt Impulsivity Scale; LOC = Locus of Control; MMPI = Minnesota Multiphasic Personality Inventory; SILS = Shipley Institute of Living Scale; STAI = State-Trait Anxiety Inventory; TAF = Thought Action Fusion Scale. **Boldface** indicates significant values.

^a*p* value represents results of χ^2 test. *ns*: *p* > .05.

^b*p* value represents result of Fischer's exact test.

^c*p* value represents results from Satterthwaite method for unequal variances.

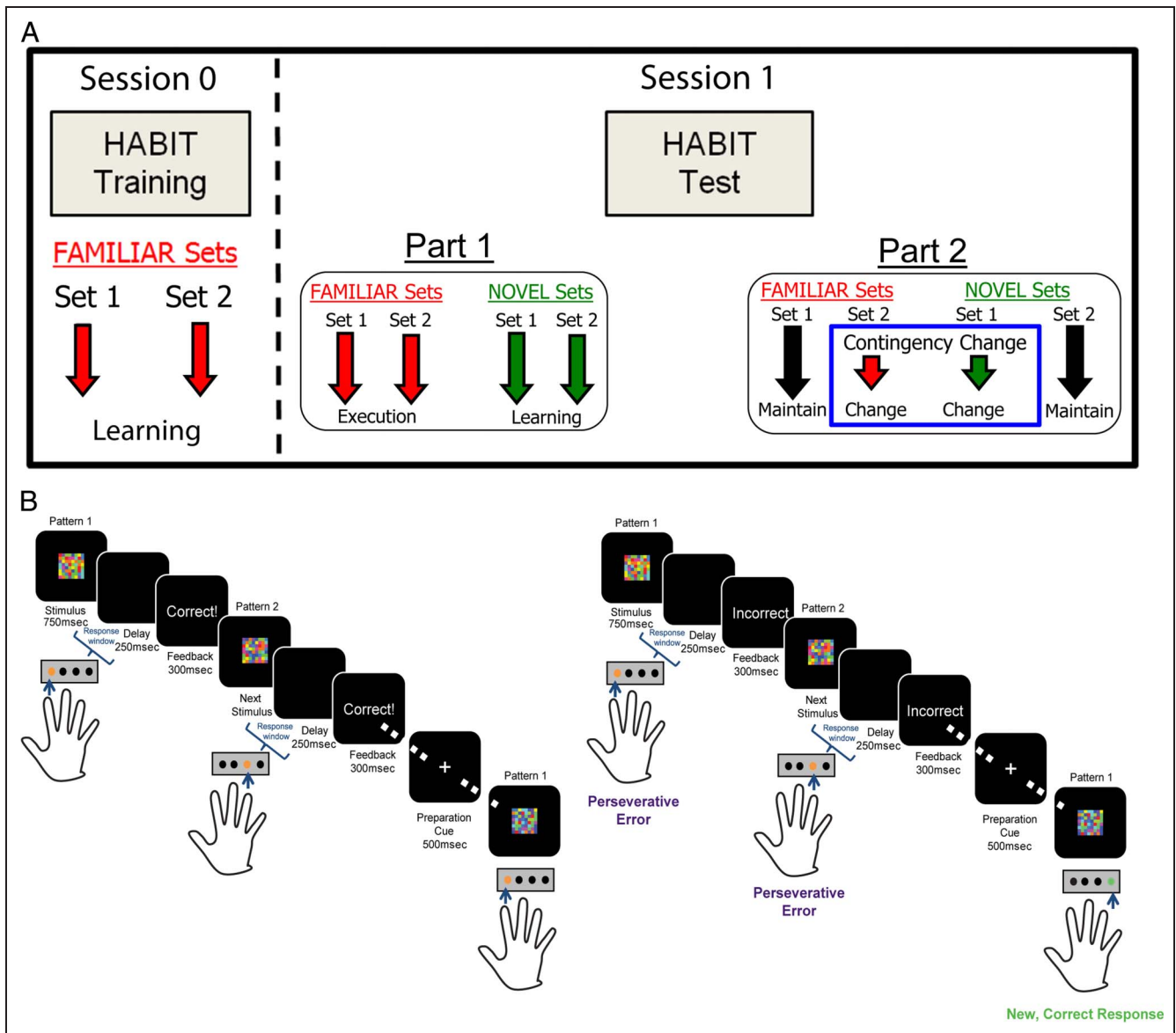


Figure 1. Diagram of HABIT paradigm structure. (A) Panel depicts training session (“Session 0”) and test session (“Session 1”), which occurs on a subsequent day. Session 1 is divided into Part 1 (preresponse change; six runs) and Part 2 (postresponse change; six runs). (B) Task schematic for Part 1 (left) and Part 2 (right) of the HABIT test session.

performance bonuses in the second (testing) session. During Session 0, participants first underwent a structured clinical interview, then completed a battery of standard questionnaires (see “Behavioral Inventories”), followed by behavioral training on the computerized S-R learning task (see “Behavioral Task”). Learning and habitual responding was then tested during Session 1.

Behavioral Inventories

We administered a number of standard questionnaires to quantify factors that could impact our results. We quantified alcohol use behavior with the Alcohol Use and Disorders Identification Test (Saunders, Aasland, Babor,

de la Fuente, & Grant, 1993) and substance use behavior with the Drug Use Screening Inventory, Domain I (Tarter, 1990) and the Drug Abuse Screening Test (Skinner, 1982). We calculated density of familial alcohol abuse using the Family Tree Questionnaire (Mann, Sobell, Sobell, & Pavan, 1985). Neuropsychological questionnaires included the Barratt Impulsivity Scale (Barratt, 1994), the Beck Depression Inventory (Beck & Steer, 1987), Rotter’s Locus of Control Scale (Rotter, 1966), the State-Trait Anxiety Inventory (Spielberger, 1985), the Thought Action Fusion Scale (Shafran, Thordarson, & Rachman, 1996), and the Anti-social Practices of the Minnesota Multiphasic Personality Inventory 2 (Butcher, Graham, Williams, & Ben-Porath, 1990). Education and occupation were quantified with

the Hollingshead Socioeconomic Status score (Hollingshead, 1975). We estimated IQ with the Shipley Institute of Living Scale (Zachary, 1991).

Behavioral Task

The HABIT is an S-R learning and relearning task implemented in E-Prime 2.0 (PST, Inc., Pittsburgh, PA) consisting of a HABIT Training Session and a two-part HABIT Test Session, which occurs on a subsequent day (Figure 1). The training and test session Part 1 have been described in detail (Boettiger & D'Esposito, 2005). In brief, stimuli were presented on a color LCD screen, and participants used a four-button keypad for manual response selection using the fingers of their dominant hand. Participants were given instructions and a brief familiarization before completing the training phase of the task. Participants viewed abstract visual stimuli displayed briefly (700 msec) on the screen that they learned, through trial and error, to associate with specific manual responses. During the first training session, participants learned two sets of S-R rules (Familiar) to a criterion of $\geq 90\%$ accuracy. Participants then returned after ≥ 1 night's sleep to complete the test session. In the second testing session, participants first demonstrated retention of the previously learned (Familiar) associations, then the learning task (HABIT Test Part 1; Figure 1) began. In the learning task, blocks of the two Familiar sets were interspersed with blocks composed of two new (Novel) stimulus sets to measure new S-R learning and blocks of a control condition, consisting of novel, unrelated stimuli (No Rule); blocks consisted of 15 randomly selected stimuli from the relevant set. Following six "runs" of 15 blocks each (three per set type), participants were informed that the correct responses for two sets (one Familiar and one Novel set) had changed (HABIT Test Part 2; Figure 1). As the previously correct responses for the changed sets produce a negative rather than positive outcome, one could construe this change in response contingency as a response "devaluation," although this manipulation is quite different from the outcome devaluation procedures traditionally used in studies of habitual responding. Devaluing outcomes is methodologically tricky in human studies, as primary rewards are not universally palatable. Moreover, points (or other performance metrics) or money tends to remain intrinsically rewarding and is difficult to realistically devalue. Participants then learned the new correct S-R associations through trial and error. This "response devaluation" manipulation allows us to quantify habitual responding when attempting to overcome both well-learned (Familiar) and freshly learned (Novel) S-R associations, as the proportion of perseverative errors can be taken as an index of the degree to which responses are outcome independent (i.e., habit-based), as opposed to outcome driven (i.e., goal-directed). By introducing S-R changes for both Familiar and Novel sets, at a point where performance is approximately equivalent, we can rule out performance deficits

due to impaired response inhibition. Moreover, including Familiar and Novel sets in which correct responses do not change allows us to control for effects on performance of time and of context change.

Data Analysis

Our main index of performance was number of correct responses out of total responses across both epochs of the task (six runs each, pre- and post-contingency change). Our data structure is composed of 48 repeated measures, consisting of four stimulus set types (2 Familiar, 2 Novel) that are measured within person over the 12 time points. We also collected RT data in each trial and were able to categorize error types (perseverative button press, other incorrect button press) postcontingency change to distinguish between habit-based and goal-directed response strategies. Because of the non-normal nature of these data, rather than using a mixed model repeated-measures ANOVA analytical approach, we instead used a GLMM with a binomial distribution and logit link function, which models a linear rate of learning and is ideally suited to account for the nonlinear nature of learning rates in terms of probability. Our GLMM approach is described in detail in the next section. To test the significance of between-group comparisons for demographic and psychological variables, we used unpaired two-tailed *t* tests for continuous measures and χ^2 tests for categorical measures. Additionally, we used a one-way ANOVA to test for statistically significant differences between groups in perseverative responding and the nonparametric Kruskal–Wallis test to compare perseverative responding among SUD subgroups. All analyses included age and IQ as covariates. All data analyses were performed within SAS (Cary, NC).

Specification of GLMMs

Performance data in this S-R learning task were the number of correct responses within each block, a nonnormally distributed outcome heavily skewed toward the top end of possible values; hence, performance accuracy was characterized by fitting GLMMs with a binomial distribution and logit link function. GLMMs provide a statistically efficient way to independently account for variance at different levels within nested data. In the present instance, repeated measures (performance within six runs each during the learning and relearning epochs) are nested within persons, and thus, GLMMs could be specified to account for both within- and between-person variability in performance accuracy. Our data structure is composed of 48 repeated measures, consisting of four stimulus set types (2 Familiar, 2 Novel) that are measured within person over the 12 time points. The measurement of accuracy at multiple time points both pre- and post-contingency change yields increased power to detect between-subject differences in within-subject change, with

particular emphasis on the ability to compare pre- and postcontingency change trajectories and to capture changes in performance over time. For each set type, we modeled the time course of performance over each epoch (pre- and post-contingency change) independently and also included additional unique variables capturing the change in performance following the S-R contingency change manipulation. Our analytic approach involved first fitting a baseline model (Model 1) to represent changes in performance accuracy during the learning and relearning epochs as a function of set type and changed response contingencies, controlling for age and IQ. Next, we added SUD status as a predictor of performance (Model 2) and conducted a likelihood ratio test to evaluate improvement in model fit. Models were estimated using maximum likelihood in the GLIMMIX procedure of SAS 9.3, implemented using adaptive quadrature with nine quadrature points per dimension of integration.

Defining π_{ij} to be the probability that person j will produce an accurate response to a stimulus given during run i , Model 1 was specified as

$$\begin{aligned} \logit(\pi_{ij}) = & \beta_{0ij} + \beta_{1ij}Trend1_{ij} + \beta_{2ij}Dropoff_{ij} \\ & + \beta_{3ij}ChangeTrend_{ij} + \beta_4Age_j + \beta_5IQ_j \\ & + u_{0j} + u_{1j}Trend1_{ij} + u_{2j}Dropoff_{ij} \\ & + u_{3j}ChangeTrend_{ij} \end{aligned} \quad (1)$$

where fixed effects are designated by β and the first four fixed effects, which capture changes in performance accuracy over time and across conditions, are decomposed as follows:

$$\begin{aligned} \beta_{0ij} &= \beta_{00} + \beta_{01}Set_{ij} \\ \beta_{1ij} &= \beta_{10} + \beta_{11}Set_{ij} \\ \beta_{2ij} &= \beta_{20} + \beta_{21}Set_{ij} + \beta_{22}NewResponse_{ij} \\ &+ \beta_{23}Set_{ij} \times NewResponse_{ij} \\ \beta_{3ij} &= \beta_{30} + \beta_{31}Set_{ij} + \beta_{32}NewResponse_{ij} + \beta_{33}Set_{ij} \\ &\times NewResponse_{ij} \end{aligned} \quad (2)$$

Last, the random effects are designated by u and assumed to be normally distributed with a full covariance matrix \mathbf{G} .

The variables within the model were coded to enhance interpretation of the parameter estimates. The covariates *Age* and *IQ* were mean-centered so that all fixed effects could be interpreted to represent effects for a participant of typical Age and IQ. *Trend1* was coded $-5, -4, \dots, 6$ for the 12 runs, *Dropoff* was coded 0 for runs occurring during the learning epoch and 1 for runs during the relearning epoch, and *ChangeTrend* was coded 0 for runs occurring during the learning epoch and 1, 2, $\dots, 6$ for runs during the relearning epoch. Given this coding, β_{0ij} represents performance at the final run of the learning epoch, β_{1ij} represents the increase in accuracy over the learning epoch, β_{2ij} represents the drop off in accuracy between the last run of the learning epoch and the first run of the relearning epoch due to changes in S-R contingen-

cies, and β_{3ij} indicates the difference in rate of improvement in accuracy in the relearning epoch relative to the learning epoch. Equation 2 shows that the values of these four coefficients were a function of set type (*Set*; coded 1 for Familiar, 0 for Novel) and whether the response was devalued in the relearning epoch (*NewResponse*; coded 1 for devalued sets and 0 for nondevalued sets). Additionally, the random effects in Equation 1 allowed for person-to-person variability in the four components of the performance accuracy trajectories.

Model 2 retains Equation 1 but includes SUD status as a predictor such that Equation 2 is elaborated as follows:

$$\begin{aligned} \beta_{0ij} &= \beta_{00} + \beta_{01}Set_{ij} + \beta_{02}SUD_j + \beta_{03}Set_{ij} \times SUD_j \\ \beta_{1ij} &= \beta_{10} + \beta_{11}Set_{ij} + \beta_{12}SUD_j + \beta_{13}Set_{ij} \times SUD_j \\ \beta_{2ij} &= \beta_{20} + \beta_{21}Set_{ij} + \beta_{22}NewResponse_{ij} + \beta_{23}Set_{ij} \\ &\times NewResponse_{ij} + \beta_{24}SUD_j + \beta_{25}Set_{ij} \times SUD_j \\ &+ \beta_{26}NewResponse_{ij} \times SUD_j + \beta_{27}Set_{ij} \\ &\times NewResponse_{ij} \times SUD_j \\ \beta_{3ij} &= \beta_{30} + \beta_{31}Set_{ij} + \beta_{32}NewResponse_{ij} \\ &+ \beta_{33}Set_{ij} \times NewResponse_{ij} + \beta_{34}SUD_j \\ &+ \beta_{35}Set_{ij} \times SUD_j + \beta_{36}NewResponse_{ij} \\ &\times SUD_j + \beta_{37}Set_{ij} \times NewResponse_{ij} \times SUD_j \end{aligned} \quad (3)$$

Models 1 and 2 are nested in their fixed effects, permitting a likelihood ratio test of the overall effect of SUD status on performance accuracy trajectories (Table 2).

RESULTS

Participants learned two sets of (Familiar) S-R associations during an initial HABIT training session, then returned for a HABIT testing session broken into two epochs: an initial learning epoch in which participants both executed the previously learned (Familiar) S-R associations and learned two new (Novel) sets of S-R associations, and a subsequent “relearning” epoch, in which the established S-R contingencies for one of the Familiar S-R sets and one of the Novel S-R sets changed (Figure 1). During the relearning epoch, the previously correct response to the stimuli in the changed sets is met with a punishment instead of a reward, reducing the value of selecting the previously learned action in response to those stimuli; this change in response contingency allowed us to quantify perseverative errors as an index of habitual responding. The learning and relearning epochs were each divided into six segments, each in turn consisted of three randomly ordered blocks of S-R set types (18 trials per block). Thus, at the onset of the relearning epoch, participants have completed 324 trials for each of the Novel S-R sets and approximately three times as many trials for the Familiar sets (average: 1038 trials; 95% CI [952, 1124]).

Table 2. Fixed Effect Estimates (Top) and Variance–Covariance Estimates (Bottom) for Models of the Predictors of Learning Behavior

Parameter	Model 1	Model 2
<i>Fixed Effects</i>		
Intercept	0.95** (0.08)	0.91** (0.10)
Set	0.33** (0.03)	0.24** (0.04)
Trend1	0.23** (0.02)	0.23** (0.02)
Set × Trend1	−0.12** (0.01)	−0.12** (0.01)
Dropoff	−0.53** (0.07)	−0.48** (0.08)
Set × Dropoff	−0.01 (0.06)	0.02 (0.07)
Dropoff × NewResponse	−0.66** (0.04)	−0.69** (0.05)
Set × Dropoff × NewResponse	0.33** (0.06)	0.39** (0.07)
ChangeTrend	−0.13** (0.02)	−0.15** (0.02)
Set × ChangeTrend	0.07** (0.02)	0.10** (0.02)
NewResponse × ChangeTrend	0.11** (0.01)	0.15** (0.02)
Set × NewResponse × ChangeTrend	−0.01 (0.02)	−0.03 (0.03)
Age (centered)	0.01 (0.01)	0.01 (0.01)
IQ (centered)	0.002 (0.01)	0.01 (0.01)
SUD		0.13 (0.18)
Set × SUD		0.28** (0.07)
Trend1 × SUD		0.02 (0.03)
Set × Trend1 × SUD		0.01 (0.02)
Dropoff × SUD		−0.14 (0.13)
Set × Dropoff × SUD		−0.11 (0.13)
Dropoff × NewResponse × SUD		0.10 (0.09)
Set × Dropoff × NewResponse × SUD		−0.17 (0.13)
ChangeTrend × SUD		0.04 (0.04)
Set × ChangeTrend × SUD		−0.09* (0.04)
NewResponse × ChangeTrend × SUD		−0.09* (0.03)
Set × NewResponse × ChangeTrend × SUD		0.08 (0.04)
<i>Variance of Random Effects</i>		
Intercept	0.40	0.38
Trend1	0.01	0.01
Changetrend	0.01	0.01
Dropoff	0.16	0.15

Table 2. (continued)

Parameter	Model 1	Model 2
<i>Correlations between Random Effects</i>		
Trend1/intercept	0.05	0.05
Changetrend/intercept	−0.04	−0.04
Changetrend/trend1	−0.01	−0.01
Dropoff/intercept	−0.14	−0.13
Dropoff/trend1	−0.03	−0.02
Dropoff/changetrend	0.01	0.01
−2*log-likelihood	23,279.02	23,207.43**

Standard errors are in parentheses. Set denotes the familiar versus novel set type variable, with novel set type as the reference category. Trend1 indicates the slope of performance during precontingency change time points. Dropoff signifies the difference in performance pre- and post-contingency change. NewResponse indicates a change in the correct response as a result of devaluation. Changetrend is the variable denoting the change in postchange performance relative to pre-change performance. The random parameters represent the variance and covariance estimates generated from inclusion of random effects in the model. The −2log-likelihood demonstrates the value for model fit.

* $p < .05$.

** $p < .001$.

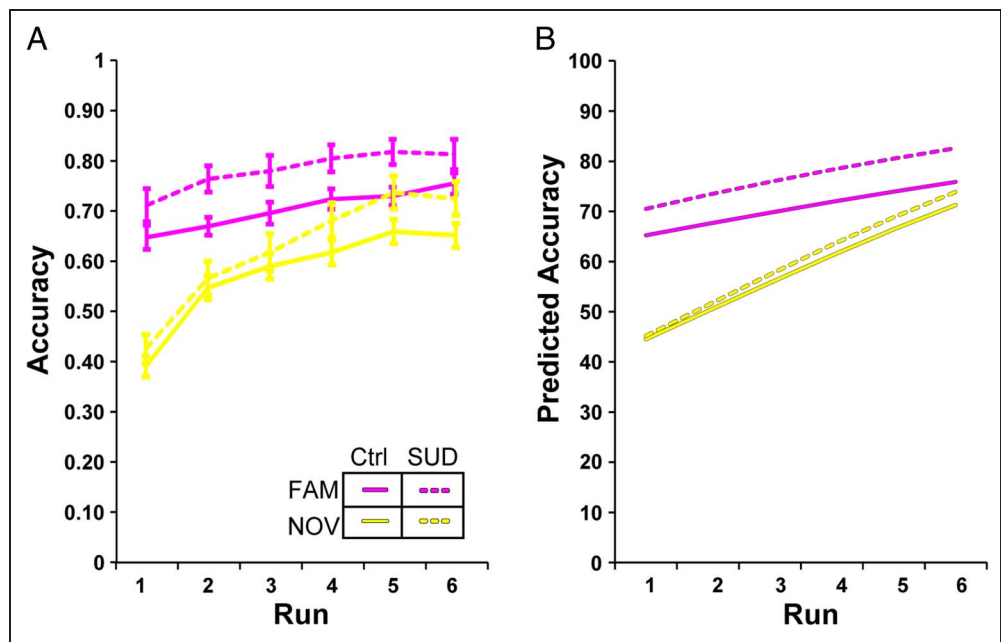
Behavioral Performance during Training Session

During the initial training session, participants were required to reach a performance criterion of 90% accuracy for each S-R set. These sets are designated as Familiar in the subsequent learning and relearning epochs. Set order was counterbalanced across participants, and set order did not differ between groups, $\chi^2_{(1)} = 0.40, p = .53$. Training to criterion took an average of ~25 min, with no significant difference between groups in the number of blocks to criterion (Ctrl: 11 blocks; SUD: 9 blocks; 40 trials per block; $F(3, 56) = 0.67, p = .57$). Learning the associative rules for the second S-R set was always more rapid and also did not differ significantly between groups (Ctrl: 4 blocks; SUD: 4 blocks; $F(3, 56) = 0.39, p = .76$). Thus, before returning for the testing session, training performance between groups was equivalent. Moreover, the time between the training and testing sessions did not differ significantly between groups (Ctrl: 10 days; SUD: 8 days; $t(60) = 1.09, p = .28$).

Behavioral Performance during Testing Session—Model 1: Baseline Model without SUD Status

We found no significant main effects of age or IQ in the baseline GLMM fit to the performance data (Table 2). During the learning epoch, we observed a significant interaction between set type and time before S-R contingency changes ($p < .001$; “Set × Trend1,” Table 2), indicating that, as expected, performance improved more over time

Figure 2. Mean accuracy values during prechange task performance. Solid lines represent the control group (Ctrl), and dashed lines represent the SUD history group. Mixed models demonstrated that group status significantly predicts accuracy (Table 2). Familiar (magenta) performance starts high and remains high as performance progresses, with SUD history predicting more accurate execution of S-R associations (Model 2, Table 2). Novel (yellow) set performance improves over time as S-R associations are learned, with no difference between groups in rate of learning (Table 2). (A) Data plots depict raw accuracy values adjusted for Age and IQ; error bars represent within-subject SEM. (B) Corresponding model predicted values.



in the Novel S-R sets relative to the Familiar S-R sets. Somewhat surprisingly, the performance drop-off effect after contingency change was greater for Novel S-R sets with changed responses contingencies relative to Familiar S-R sets with changed responses contingencies ($p < .001$; “Set \times Dropoff \times NewResponse,” Table 2).

During the relearning epochs the difference in learning rate interacted with S-R set type, with unchanged Novel S-R sets showing a shallower rate of improvement ($p < .001$; “ChangeTrend,” Table 2), which was less pronounced for unchanged Familiar S-R sets ($p < .001$; “Set \times ChangeTrend,” Table 2). This difference in learning rate between epochs also differed between S-R sets with changed versus unchanged response contingencies ($p < .001$; “NewResponse \times ChangeTrend,” Table 2). For changed S-R sets, the rate of relearning was steeper than that observed during the learning epoch, in contrast to the shallower relearning rate for unchanged sets.

Model 2: Including SUD Status

Across both epochs, a model including group as a performance predictor (Model 2, Table 2) fit the data significantly better than did an identical model excluding group as a predictor (Model 1, Table 2; $p < .001$). This result indicates that presence or absence of an SUD history accounted for significant variability in HABIT performance across individuals. As described below, this result does not reflect a performance deficit in the SUD group.

To further dissect HABIT performance, we first evaluated the initial learning epoch. Task performance improved over the course of the epoch, with greater improvement for the Novel S-R sets (Figure 2; $p < .001$, “Set \times Trend1,”

Table 2). As shown in Figure 2, participants executed Familiar S-R sets more accurately than Novel S-R sets, a distinction that was heightened in the SUD group. The groups did not differ in terms of performance improvement during the initial learning epoch ($p = .50$; “Set \times Trend1 \times SUD,” Table 2), but an SUD history predicted more accurate execution of Familiar S-R sets (Figure 2, magenta lines; $p < .001$; “Set \times SUD,” Table 2). Thus, an SUD history predicts intact ability to form new S-R associations and a somewhat heightened ability to accurately execute established S-R associations.

At the outset of the relearning epoch, performance immediately declined for all sets in both groups as shown in Figure 3 (right). As in Model 1, the changed–unchanged S-R contingency contrast was more pronounced in the Novel S-R sets (yellow) relative to the Familiar S-R sets (magenta; $p < .001$; “Set \times Dropoff \times NewResponse,” Table 2); SUD status did not significantly interact with these parameters (Table 2). This finding indicates that both groups show evidence of overtraining in the Familiar S-R sets relative to the Novel S-R sets, which is reported to facilitate reversal learning for S-R tasks (McLaren et al., 2014). This is consistent with the fact that participants completed two to three times as many trials for the Familiar S-R sets relative to the Novel S-R sets ($n = 324$ trials per set). As is evident in Figure 3, performance improved over the course of the relearning epoch, with shallower rates of increase relative to the initial learning epoch for unchanged, Novel sets ($p < .001$; “ChangeTrend,” Table 2), an effect that did not differ by group ($p = .35$; “ChangeTrend \times SUD,” Table 2). For Novel sets with changed S-R contingencies, control participants demonstrated steeper rates of performance improvement postchange

relative to prechange ($p < .001$; “NewResponse \times Change Trend,” Table 2). In contrast, the SUD group showed a shallower rate of improving performance for Novel changed sets postchange relative to prechange. For Familiar unchanged sets, the control group demonstrated a

steeper rate of improvement in the relearning epoch relative to the prechange epoch ($p < .001$; “Set \times Change Trend,” Table 2). During relearning, the SUD group showed significantly shallower rates of performance improvement for unchanged Familiar S-R sets ($p < .05$;

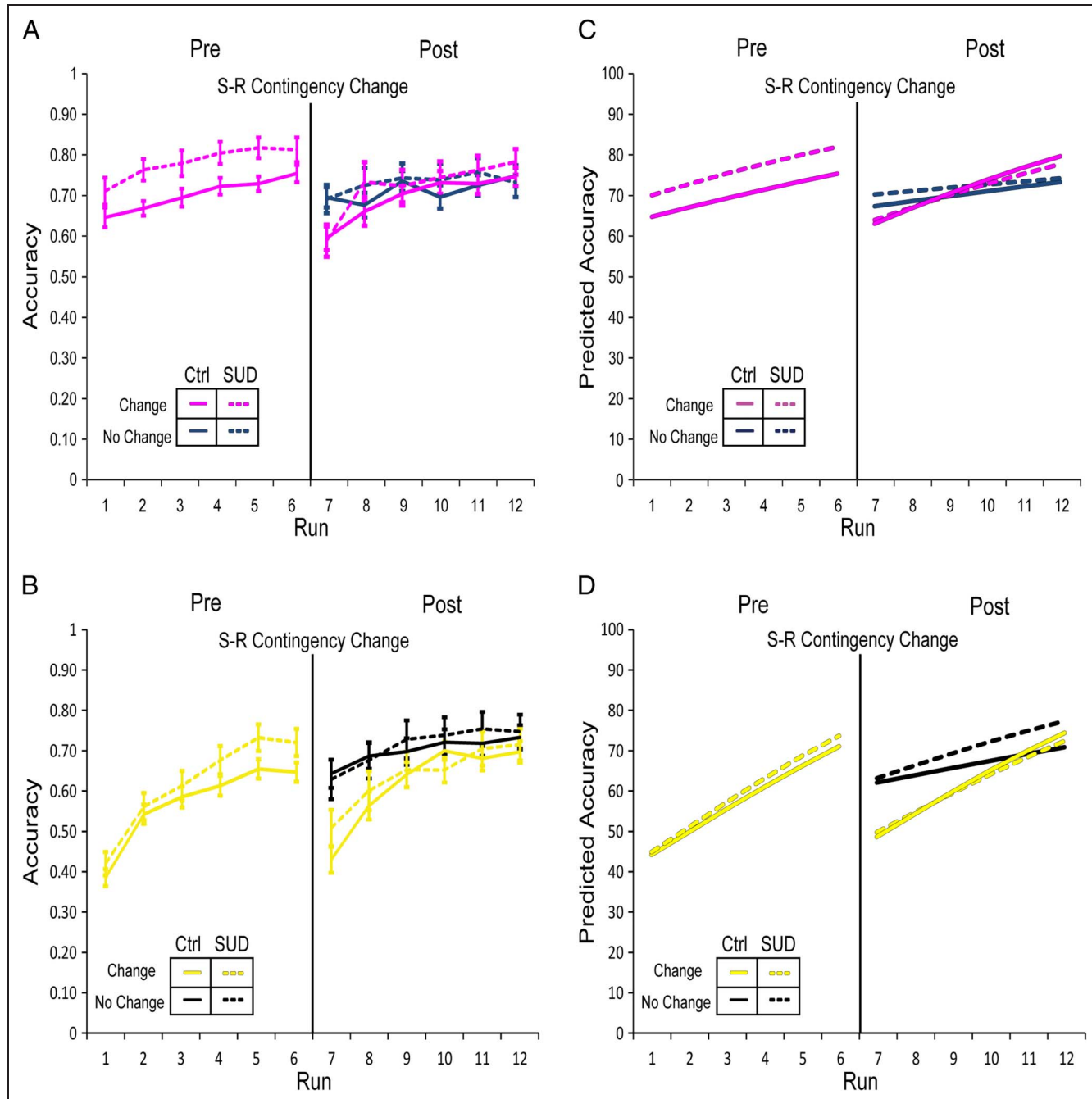


Figure 3. Mean accuracy values during pre- and postchange task performance. Solid lines represent the control group and dashed lines represent the SUD history group. Overall, a dropoff in performance occurred in blocks with changed response contingencies and performance improved over the course of relearning (“Post” panels; Table 2); the rate of relearning compared with the initial (“Pre”) learning rate was dependent on group status, set type, and change status (Table 2). (A) Panels depict performance in Familiar sets (magenta) during pre- and postchange. Dark blue lines in the right panel (“Post”) indicate performance in the set with unchanged response contingencies during the relearning phase; performance in the response-changed set shown in magenta. (B) Performance in Novel sets (yellow) during pre- and postchange. Black lines in the right panel (“Post”) indicate performance in the set with unchanged response contingencies during relearning; performance in the response-changed set shown in yellow. Performance dropped more dramatically after response change for Novel sets relative to Familiar sets (Table 2). Additionally, the control group showed steeper learning rates for Novel changed sets relative to the SUD group. Corresponding model predicted values are shown in panels C and D.

“Set × ChangeTrend × SUD,” Table 2), whereas the interaction between SUD status, set type, contingency change, and the change in learning rate in the relearning epoch was not significant ($p = .08$; “Set × NewResponse × ChangeTrend × SUD,” Table 2). To summarize, S-R contingency change did not reveal a global impairment in response flexibility or inhibitory control among people with an SUD history. In fact, for the changed Novel S-R set, the SUD group’s performance was less impaired than that of the control group immediately following contingency change, resulting in a more rapid performance recovery for the SUD group (Figure 3, yellow dashed line).

Responding to a stimulus with an action that is no longer valued (i.e., no longer positively reinforced) is taken as an indicator of habit-based, rather than goal-directed, responding. Thus, to quantify the habitual nature of responding after response contingency change, we evaluated the percentage of perseverative errors during the relearning epoch. A one-way ANOVA between group for each set type indicated significant differences between groups for the overall percentage of perseverative errors for the Familiar set type ($p = .004$), but not for the Novel set type ($p = .43$). These results reflect the fact that when trying to replace the well-established Familiar S-R associations, errors made by the SUD group were more apt to be perseverative errors ($p = .002$; Figure 4). No

such group difference was observed for replacement of more recently established Novel S-R associations ($p = .146$; Figure 4). These findings indicate the more habitual nature of responding in the Familiar S-R sets among SUD participants.

To evaluate the contribution of abused substance type to perseverative responding during S-R relearning, we stratified SUD participants into two categories: history of stimulant dependence ($n = 11$) or no history of stimulant dependence ($n = 11$). We found a significant difference in perseverative errors during Familiar S-R relearning (Kruskal–Wallis test, $p = .009$; Figure 5, pink bars), but not during Novel S-R relearning ($p = .182$; Figure 5, yellow bars). Post hoc tests (Bonferroni corrected, $p < .025$) demonstrated that participants with or without a history of stimulant dependence made significantly more perseverative errors during relearning of Familiar S-R sets relative to controls (stimulant history, $p = .007$; no stimulant history, $p = .009$). However, only the stimulant dependence group showed a trend toward more perseverative responding during Novel S-R relearning relative to controls (stimulants; $p = .038$; no stimulants; $p = .342$). The results in the Novel condition suggest that stimulant addiction may be associated with an even more rapid transition to habitual responding.

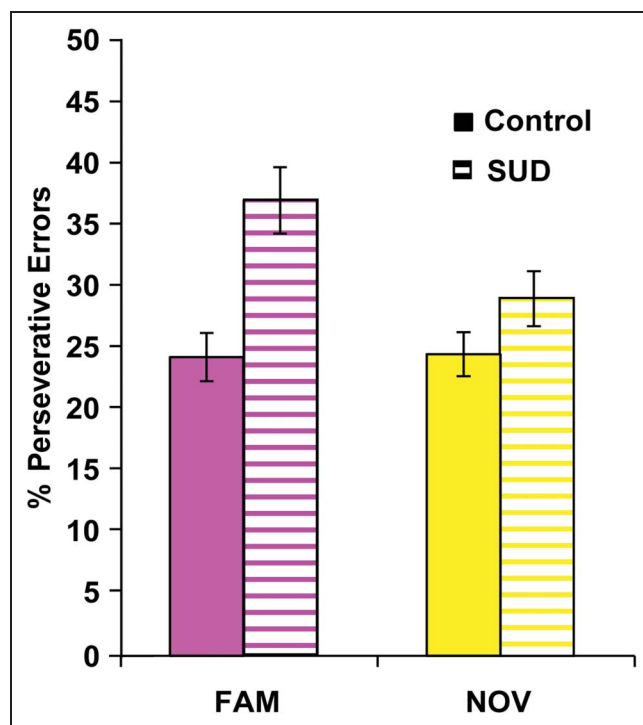


Figure 4. Postchange percentage of perseverative errors by group for S-R sets. The percentage of perseverative errors in the Familiar set (magenta) with changed response contingencies significantly differed by group, with SUD history participants making incorrect responses that were perseverative in nature, $F(3, 58) = 4.88, p = .004$. In contrast, the percentage of perseverative errors in the Novel set (yellow) with changed response contingencies did not differ between groups, $F(3, 58) = 0.93, p = .43$. Error bars represent SEM.

DISCUSSION

We demonstrate that people with SUDs learn new S-R associations as well as control participants do and can flexibly adapt newly learned S-R associations but are specifically impaired in overcoming well-learned S-R associations. Notably, those with SUDs differ from controls only in terms of perseverative errors committed during well-established S-R replacement, indicating the more habit-based nature of their responses. These findings suggest that people with a history of an SUD more rapidly acquire habitual responding outside the drug-taking domain.

Prior Studies Linking Habit and Addiction

Despite extensive investigation of drugs of abuse and habit in animal models, modest translation of these experimental paradigms to human studies has occurred to date. Young, light-smoking adults will pursue both cigarette and chocolate rewards via goal-directed strategies (Hogarth & Chase, 2011). Nicotine dependence was low in this sample; however, Hogarth, Chase, et al. (2012) made similar findings in a sample of daily and nondaily smokers, in addition to finding a positive correlation between motor impulsiveness and habitual responding. These studies suggest that habitual drug consumption may associate with personality factors that predispose individuals toward habit-based responding.

Hogarth and colleagues have also found that acute alcohol intake renders the selection strategy for both water and chocolate rewards habitual (Hogarth, Attwood,

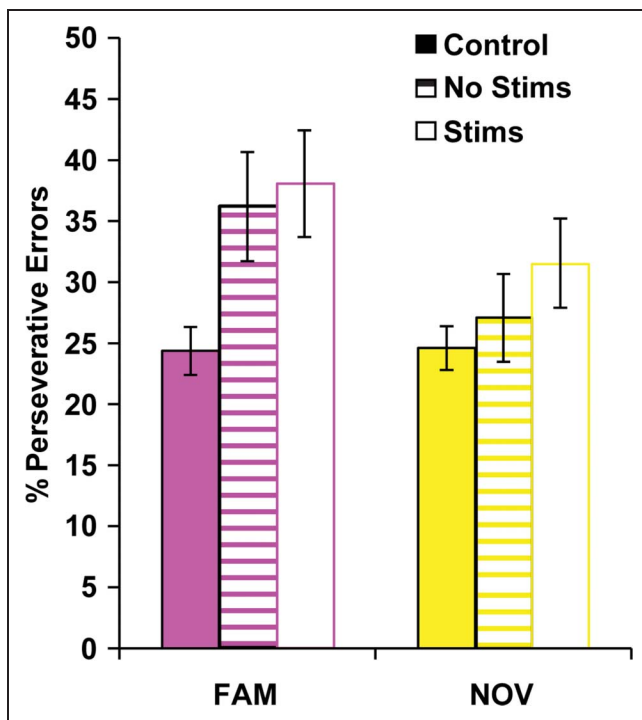


Figure 5. Postchange percentage of perseverative errors by abused substance type group for S-R sets. We categorized participants according to substance dependence history as follows: no history (control), no stimulant dependence (alcohol or opiate dependence; “No Stims”), or stimulant and alcohol dependence (stimulants; “Stims”). Overall nonparametric comparison of the three groups indicated a significant difference in the percentage of perseverative errors for the Familiar set (magenta), $\chi^2(3) = 11.67, p = .009$. The groups did not differ in terms of the percentage of perseverative errors for the Novel set (yellow), $\chi^2(3) = 4.86, p = .182$. Post hoc tests corrected for multiple comparisons ($p = .025$) demonstrated that, compared with controls, participants with a history of stimulant dependence committed a higher percentage of perseverative errors for both the Familiar set ($z = 2.69, p = .007$) and the Novel set ($z = 2.07, p = .038$). In contrast, participants with no history of stimulant dependence committed a higher percentage of perseverative errors when compared with controls only in the Familiar set ($z = 2.60, p = .009$), not in the Novel set ($z = 2.69, p = .342$).

Bate, & Munafo, 2012), which is consistent with data showing that exposure to alcohol potentiates habitual responding in rats (Corbit et al., 2012). A recent fMRI study of alcohol-dependent patients in an instrumental choice task (de Wit, Niry, Wariyar, Aitken, & Dickinson, 2007) found evidence of preferential S-R based responding, rather than goal-directed actions, in alcohol-dependent individuals relative to controls (Sjoerds et al., 2013). Furthermore, relative to controls, the alcohol-dependent group increased activation of the posterior putamen and reduced vmPFC activation during instrumental choice. Although these data are consistent with the animal literature associating chronic exposure to drugs of abuse with an overreliance on S-R response strategies, many of the alcohol-dependent patients in Sjoerds et al.’s study were concurrently using psychoactive medications for depression and anxiety disorders, precluding unequivocal attribution of group differences solely to alcohol dependence.

Regardless, these results point to neural correlates within frontostriatal circuits for enhanced reliance on S-R versus goal-directed actions. Such findings enable strong predictions about expected differences between people with SUDs and healthy controls in terms of neural activation associated with response selection in the HABIT.

Neurobiology of Habit in Humans

Although the neural bases of behavioral differences in S-R learning among people with SUDs is largely unexplored, the SUD neuroimaging literature suggests that alterations in frontostriatal circuit recruitment underlie atypical behavior in SUDs (Ersche et al., 2012; Konova et al., 2012; Goldstein & Volkow, 2011; Koob & Volkow, 2010; Park et al., 2010; Kalivas, 2008; Olausson et al., 2007). Although this work has not investigated habits per se, it logically follows that impaired frontal control of striatal output signals could yield overreliance on striatal habit circuits. Drugs of abuse may alter frontostriatal circuitry (Izquierdo & Jentsch, 2012) such that frontal input to the striatum can no longer effectively act as a “switch” to regulate the contribution of brain signals for automatic, habitual responding versus goal-directed action (Smith & Graybiel, 2013; Smith, Virkud, Deisseroth, & Graybiel, 2012).

A unique aspect of the HABIT paradigm is the ability to measure behavior during attempts to overcome habitual responding; using perseverative errors as an index of habit-based responding and continuing task conditions that require goal-directed responding allows us to measure the flexibility of behavior over an extended period after response contingency change. Essentially, we are able to measure the ability to “break” habits that have been formed within this task. The ability to change or “break” habitual behaviors has not directly been tested in humans to our knowledge, but converging evidence from both recent animal and human studies relate drugs of abuse and frontostriatal circuitry to the regulation of behavioral change. In primates, prolonged cocaine intake profoundly impairs S-R relearning (Jentsch, Olausson, De la Garza, & Taylor, 2002). These data suggest that chronic drug exposure potentiates habitual response selection and further supports a role for extended substance abuse in altering the circuits underlying S-R learning and replacement. In rodents, optogenetic perturbation of the infralimbic portion of the mPFC results in a switch from a previously to recently learned behavior and thus facilitates the replacement of habitual behaviors (Smith et al., 2012). Computational modeling of human choice behavior in which prefrontal brain regions “arbitrate” between habit-based or goal-directed responses further supports these animal findings (Lee, Shimojo, & O’Doherty, 2014). TMS applied to the DLPFC shifts the balance between goal-directed versus habit-based response selection strategies (Smittenaar, FitzGerald, Romei, Wright, & Dolan, 2013; Knoch, Brugger, & Regard,

2005). Taken together, these studies provide compelling evidence for the regulation of behavioral control via frontostriatal circuitry. An important future direction is to determine whether abnormal functioning of these same frontostriatal circuits underlies the atypical S-R learning and replacement we find in people with a history of addiction.

It is important to note that elevated perseverative errors in the changed Familiar S-R sets in the SUD group is unlikely to reflect impaired response inhibition, as inhibitory impairments should have manifest in the Novel condition as well as the Familiar condition based on nearly equal performance in the Novel condition before the contingency change, particularly among the SUD group. However, we only observed this deficit in the highly practiced Familiar condition in which the SUD group appears to have transitioned to a more automatic S-R strategy. One could make the case that suppressing a more automatized action requires greater response inhibition and that only under this higher “inhibitory load” condition did a deficit in the SUD group emerge; however, this argument merely lends support to our interpretation of a more rapid transition to an automatic response strategy in people with an SUD history.

Possible Role of Stress in the SUD Group Findings

Although we make the case above that atypical frontostriatal function likely underlies the apparent earlier switch to dominance of habit-based responding in the SUD group, a growing body of literature shows that stress can potentiate habit-based responding in both humans (Schwabe & Wolf, 2010, 2011, 2013; Schwabe et al., 2007) and animal models (Dias-Ferreira et al., 2009). This tendency for stress to shift the balance of behavior from goal-directed to habitual has been investigated pharmacologically in humans, with evidence indicating roles for both elevated cortisol levels and increased noradrenergic activity (Schwabe, Tegenthoff, Hoffken, & Wolf, 2012; Schwabe, Hoffken, Tegenthoff, & Wolf, 2011). Moreover, neuroimaging research has found that participants subject to chronic psychosocial stress fail to change responses after outcome devaluation, indicative of habit-based actions (Soares et al., 2012). The stressed participants in that study showed greater activation of the putamen during response selection after devaluation, relative to nonstressed controls, consistent with prior links between the putamen and habitual responding (de Wit et al., 2009; Tricomi et al., 2009). Notably, the behavioral and neural effects of stress in the Soares study were reversible, declining after the stressful period ended; this demonstrates the plasticity of the neural systems regulating habitual actions and holds promise for interventions to facilitate behavioral change of ingrained behaviors.

The evidence that stress can promote habit-based responding, together with evidence of dysregulated hypothalamic-pituitary axis function in individuals with SUDs

(Lijffijt, Hu, & Swann, 2014; Porcu, O’Buckley, Leslie Morrow, & Adinoff, 2008; Kreek, Nielsen, Butelman, & LaForge, 2005; King et al., 2002; Anthenelli, Maxwell, Geraciotti, & Hauger, 2001), suggest that the behavioral differences that we observed in the SUD group could reflect greater stress levels in the SUD group. We did not collect physiological or subjective report measures of stress for this study, although we did collect measures of anxiety; the groups did not differ in terms of state anxiety, but the SUD group did report slightly higher levels of trait anxiety (Table 1), suggesting possibly higher levels of chronic stress in the SUD group compared with controls. Stress is well known to precipitate relapse (Sinha, 2012), and although the underlying mechanisms are not well understood, it is tempting to speculate that a contributing factor could be stress-induced promotion of habitual responding in people with SUDs. This question can be addressed with the HABIT paradigm, which may ultimately identify new therapeutic approaches to relapse prevention.

Study Limitations

The observation of more habit-based responding in the SUD group could be a consequence of chronic drug exposure or a predisposing trait that contributes to SUD vulnerability; these alternatives cannot be disentangled by the current study. If this heightened propensity to establish habits in people with SUDs predates the SUD, it would represent a promising, unexplored intermediate phenotype for SUDs. A further limitation is the SUD sample studied. Participants were recruited based on any lifetime history of an SUD (including an alcohol use disorder), which yielded a heterogeneous population. Given the distinct effects of differing abused substances on neurotransmitter systems, it is rather unlikely that the behavioral effects we observed reflect common neural dysfunction caused by chronic substance abuse. However, biological predispositions play a large role in SUDs and that heritability is not necessarily substance-specific (Hicks, Iacono, & McGue, 2012). As such, one would expect to find common neural substrates across different substance abuse categories underlying shared behavioral deficits that represent preexisting vulnerability factors. Our finding here of propensity to more rapidly transition to habit-based responding could theoretically contribute to establishing and/or maintaining compulsive, habitual substance use and, as such, could theoretically play a role as a preexisting risk trait. Be that as it may, this heterogeneity, coupled with our small sample size, precludes drawing conclusions regarding specific substances or polysubstance use. The range of disease severity was also limited in our sample, with all participants falling at the severe end (range = 6–11; Hasin et al., 2013); thus, we were unable to assess whether SUD severity correlates with greater propensity for habitual responding. Another limitation was our exclusion of individuals currently using

any psychoactive medications or with comorbid mental health disorders, neurological conditions, or below normal IQ. The advantage of this “clean” sample is our confidence in attributing group differences in behavior to SUD history, but among the SUD population at large, comorbidities and psychoactive medication use is common. These exclusion criteria also likely and substantially increased our power to detect group effects, as different comorbid conditions may have either amplified or compensated for excess habitual responding; psychoactive medications may have similarly increased variance. Finally, our participants with SUDs were abstinent from substance use, and as such, their engagement of motivational circuitry and the ability to form habitual associations might be substantially different from people in the active phase of an SUD. These limitations point to key future avenues of research that will expand the scope of our understanding of habit-based responding in addiction.

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